FACSIMILE COVER SHEET

RUDEN

MCCLOSKY

SMITH

SCHUSTER &

RUSSELL, P.A. ATTORNEYS AT LAW 215 SOUTH MONROE STREET SUITE 815 TALLAHASSEE, FLORIDA 32301

TELEPHONE: (850) 681-9027 FAX: (850) 224-2032

April 20, 1998

TO:

Clair Fancy

FAX NUMBER:

922-6979

TELEPHONE NO.:

488-1344

FROM:

Mary Smallwood

CLIENT NAME:

Ogden Martin

CLIENT NO.:

29050-0002

NO. OF PAGES:

2 (INCLUDING THIS COVER PAGE)

COMMENTS:

If there are any problems or complications, please notify us immediately at (850) 681-9027.

THE INFORMATION CONTAINED IN THIS FACSIMILE MESSAGE IS ATTORNEY PRIVILEGED AND CONFIDENTIAL INFORMATION INTENDED ONLY FOR THE USE OF THE INDIVIDUAL OR ENTITY NAMED ABOVE. IF THE READER OF THIS MESSAGE IS NOT THE INTENDED RECIPIENT, YOU ARE HEREBY NOTIFIED THAT ANY DISSEMINATION, DISTRIBUTION OR COPY OF THIS COMMUNICATION IS STRICTLY PROHIBITED. IF YOU HAVE RECEIVED THIS COMMUNICATION IN ERROR, PLEASE IMMEDIATELY NOTIFY US BY TELEPHONE (IF LONG DISTANCE, PLEASE CALL COLLECT) AND RETURN THE ORIGINAL MESSAGE TO US AT THE ABOVE ADDRESS VIA THE U.S. POSTAL SERVICE. THANK YOU.

RUDEN
MCCLOSKY
SMITH
SCHUSTER &
RUSSELL, P.A.

215 SOUTH MONROE STREET SUITE 815 TALLAHASSEE, FLORIDA 32301

TELEPHONE: (850) 681-9027 FAX: (850) 224-2032

E-MAIL: MFS@RUDEN,COM

April 20, 1998

Clair Fancy
Bureau of Air Regulation
Department of Environmental Protection
2600 Blairstone Road, MS 5505
Tallahassee, FL 32399-2400

Lake County Resource Recovery Facility - Test Burn of Biomedical Waste

Dear Clair:

Re:

Thank you for meeting with me last Friday to discuss the protocol for testing of the Lake County Resource Recovery Facility with respect to biomedical waste. The purpose of this letter is to reiterate my understanding of the results of that meeting. Ogden Martin Systems of Lake, the operator of the facility, had requested permission to burn biomedical waste during a scheduled test on April 21, 1998. After reviewing the various permit conditions, it was agreed that the construction and operation permits allowed such a test burn to occur. We understand that conducting the test does not necessarily authorize the continued processing of biomedical waste at the facility.

In particular, it was agreed that biomedical waste could be tested in both Unit 1 and Unit 2 at the permit authorized throughput. The test results will be provided to the Department and maybe used in reviewing any subsequent permit modifications.

Again, Ogden appreciates your assistance in resolving this matter prior to the scheduled test date.

Sincerely,

RUDEN, McCLOSKY, SMITH, SCHUSTER & RUSSELL, P.A.

Mary F. Smallwood

MFS/cc

cc: Len Kozlov Kurt Rieke Jason Gorrie the control of the co

1	INAME	AIRS_ID	TYPE S	EU_	EU_DESCRIPTION	POLLU	TEST_ R	TEST_ALLOW	ACTUAL	UNIT	AUDI1
		_	Т	NO		TANT	DATE E	_			TYPE
			A				s			j	
			T				U				
			U				L				ĺ
1			_ S				T				
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM	0690046	3 A	2	MSW INCINERATOR #2, 288 TPD C	PM	09-Jan-97 P	0.015	0.0002	22	2
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM	0690046	3 A	2	MSW INCINERATOR #2, 288 TPD C	PM	09-Jan-97 P	0.02	0.0002	20	2
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM	0690046	3 A	2	MSW INCINERATOR #2, 288 TPD C	co ,	09-Jan-97 P	100	20.6	31	2
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM	0690046	3 A	2	MSW INCINERATOR #2, 288 TPD C	H114	30-Jun-95 P	_ 0.0003	0.000004	02	2
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM	0690046	3 A	2	MSW INCINERATOR #2, 288 TPD C	H114	30-Jun-95 P	0.0003	0.000004	02	2
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM	0690046	3 A	2	MSW INCINERATOR #2, 288 TPD C	H114	01-Feb-96 P	70	15.4	20	2
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM	0690046	3 A	2	MSW INCINERATOR #2, 288 TPD C	H114	01-Feb-96 P	0.00034	0.000007	22	2
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM	0690046	3 A	1	MSW INCINERATOR #2, 288 TPD C	ı	10-Jan-97 P	70	32.2	28	2
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM	0690046	3 A	2	MSW INCINERATOR #2, 288 TPD C	H114	10-Jan-97 P	0.00034	0.000013	22	2
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM	0690046	3 A	2	MSW INCINERATOR #2, 288 TPD C	H114	10-Jan-97 P	999990	0.00279	PH	2
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM	0690046	3 A	2	MSW INCINERATOR #2, 288 TPD C	H114	10-Jan-97 P	999990	0.024	PH	2
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM	0690046	3 A	2	MSW INCINERATOR #2, 288 TPD C	SO2	15-Jan-97 P	60	4.8		2
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM	0690046	3 A	2	MSW INCINERATOR #2, 288 TPD C	co	15-Jan-97 P	100	20.6	04	2
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM	0690046	3 A	2	MSW INCINERATOR #2, 288 TPD C	H106	15-Jan-97 P	50	28.9	04	2
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM	0690046	3 A	1	MSW INCINERATOR #2, 288 TPD C		15-Jan-97 P	0.015	0.0002	02	2
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM	0690046	3 A	2	MSW INCINERATOR #2, 288 TPD C	PM	28-Jan-98 P	0.015	0.00744	02	2
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM	0690046	3 A	2	MSW INCINERATOR #2, 288 TPD C	H106	28-Jan-98 P	50	27.1	04	2
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM	0690046	3 A	2	MSW INCINERATOR #2, 288 TPD C	NOX	28-Jan-98 P	385	322	04	2
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM	0690046	3 A	2	MSW INCINERATOR #2, 288 TPD C	co	28-Jan-98 P	100	19.2	04	2
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM	0690046	3 A	2	MSW INCINERATOR #2, 288 TPD C	SO2	28-Jan-98 P	60	16.5	04	2
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM	0690046	3 A	2	MSW INCINERATOR #2, 288 TPD C	SO2	21-Apr-98 P	60	2.2	04	2
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM	0690046	3 A	2	MSW INCINERATOR #2, 288 TPD C	H114	21-Apr-98 P	0.0003	0.000007	02	2
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM	0690046	3 A		MSW INCINERATOR #2, 288 TPD C		21-Apr-98 P	100	21.7	04	2
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM	0690046	3 A	2	MSW INCINERATOR #2, 288 TPD C	NOX	21-Apr-98 P	385	269	I	2
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM	0690046	3 A	2	MSW INCINERATOR #2, 288 TPD C	H106	21-Apr-98 P	50	11	04	2
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM		3 A		MSW INCINERATOR #2, 288 TPD C		21-Apr-98 P	0.015	0.0013		2
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM		3 A		MSW INCINERATOR #2, 288 TPD C		23-Apr-98 P	0.015	0.0013		2
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM		3 A		MSW INCINERATOR #2, 288 TPD C		23-Apr-98 P	50		04	2
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM		3 A		MSW INCINERATOR #2, 288 TPD C	,	23-Apr-98 P	385	269.3		2
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM		3 A		MSW INCINERATOR #2, 288 TPD C		23-Apr-98 P	100			2
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM		3 A		MSW INCINERATOR #2, 288 TPD C		23-Apr-98 P	60	2.2	1	2
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM		3 A		MSW INCINERATOR #2, 288 TPD C		23-Apr-98 P	0.0003	0.000007		2
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM		3 A		MSW INCINERATOR #2, 288 TPD C		29-Jan-99 P	60	0.421	1	3
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM		3 A		MSW INCINERATOR #2, 288 TPD C	l	29-Jan-99 F	70	258		3
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM		3 A		MSW INCINERATOR #2, 288 TPD C	•	29-Jan-99 P	100		04	3
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM		3 A	** * * * * * * * * * * * * * * * * * * *	MSW INCINERATOR #2, 288 TPD C	l	29-Jan-99 P	385			3
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM		3 A		MSW INCINERATOR #2, 288 TPD C		29-Jan-99 P	50	l .	04	3
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM		3 A	1	MSW INCINERATOR UNIT #1	voc	30-Jan-96 P	70	4.67		2
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM		3 A	1	MSW INCINERATOR UNIT #1	PB	31-Jan-96 P	0.00031	0.000001	1	2
OGDEN MARTIN SYSTEMS	OGDEN MARTIN SYSTEM		3 A			PB	08-Jan-97 P	0.00031		22	2 .
OGDEN MAKTIN STSTEMS	OGDEN MAKTIN STSTEM	0090046	j JA	1	INION INCINERATOR UNIT #1	120	1 00-Jan-9/ P	0.00031	1 0	22	4

POLLUTANTS & THEIR RESPECTIVE LIMITS: AS EXTRACTED FROM REGULATIONS

- 1. PARTICULATE MATTER in gases discharge limit is 27 mg/day within a large facility and 70 mg/day within a small facility
- 2. OPACITY limit is 10% (6-minute average) within a small or a large facility
- 3. CADMIUM limit is 0.040 mg/day for a larger facility and 0.10 mg/day for a small facility corrected to 7% OXYGEN.
- 4. LEAD limit is 0.49 mg/day for a large facility and 1.6 mg/day for a small facility corrected to 7% OXYGEN.
- 5. MERCURY limit is 70 micrograms per dry standard cubic meter of flue gas corrected to 7% O_2 , or 20% by weight of the mercury in the flue gas upstream of the mercury control device (80% reduction by weight), whichever occurs first.
- 6. SULFUR DIOXIDE limit is 31 ppm for a large facility and 80 ppm for a small facility corrected to $7\% O_2$.
- 7. HYDROGEN CHLORIDE limit is 31 ppm for a large facility and 250 ppm for a small facility corrected to 7% O₂
- 8. DIOXINS/FURANS limit is 60 ng for ESP based and 30 ng for non-ESP based corrected to 7% O_2 .
- 9. DIOXIN/FURANS STATE PLAN limit is 125 ng for small facility

10. CARBON MONOXIDE:

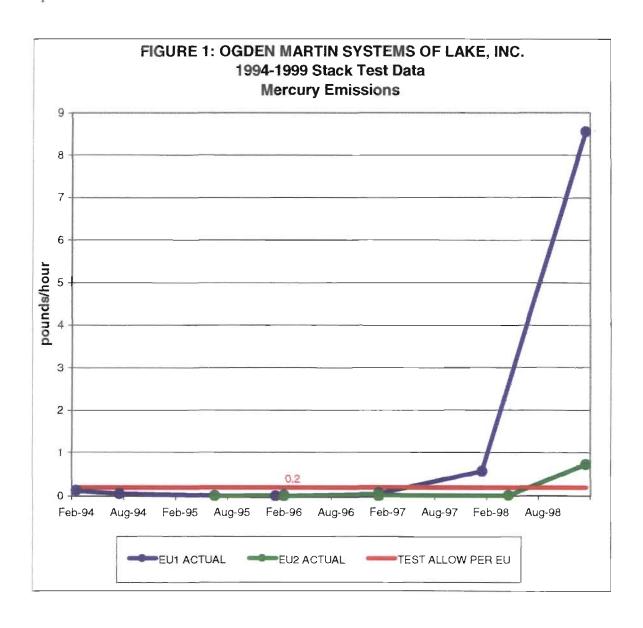
i.	Water Wall:	100 ppm	4 Hours Averaging Time
ii.	Refractory:	100 ppm	4 Hours Averaging Time
iii.	Rotary Ref.:	100 ppm	24 Hours Averaging Time
iv.	Rotary Water Wall:	250 ppm	24 Hours Averaging Time
v.	Modular Starved:	50 ppm	4 Hours Averaging Time
vi.	Modular Excess:	50 ppm	4 Hours Averaging Time
vii.	Refuse Derived Fuel:	200 ppm	24 Hours Averaging Time
viii.	Circulating Fluidized:	100 ppm	4 Hours Averaging Time
ix.	Buddling Fluidized	100 ppm	4 Hours Averaging Time
X.	Pulverized Coal/Refuse	: 150 ppm	4 Hours Averaging Time
хi.	Spreader Stoker/Refuse	: 200 ppm	24 Hours Averaging Time

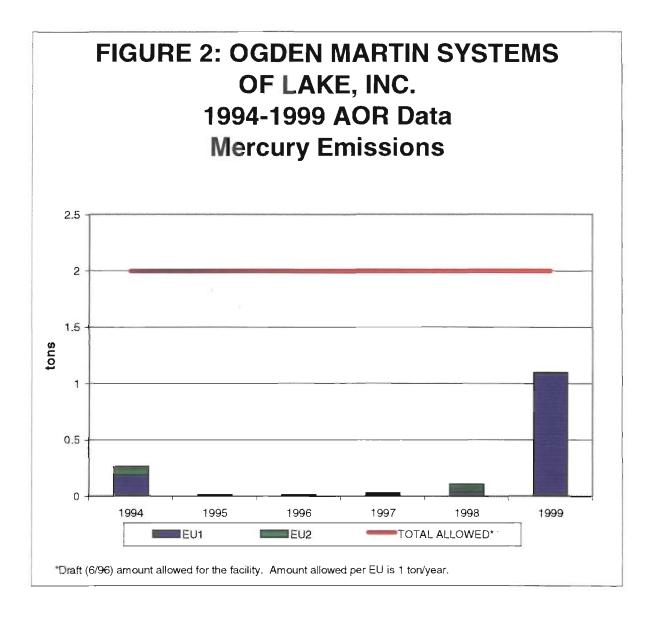
11. NITROGEN OXIDES:

1.	Water Wall:	200 ppm
ii.	Rotary Water Wall	: 250 ppm
iii.	Refuse Derived:	250 ppm
iv.	Fluidized Bed:	240 ppm
v.	Refractory:	NA
vi.	Other:	200 ppm

TREND ANALYSIS OF THE MERCURY DATA OF OGDEN MARTIN SYSTEMS OF LAKE INC

Trend analysis is used to test different aspects of the shape of a function relating a *quantitative* independent variable (permitted/allowable/permissible emission in this case) and the dependent variable (actual emissions observed/recorded) during the mercury testing. The interest is in the shape of the function relating the levels of this quantitative independent variable to the dependent variable.





Consider the Figures 1 and 2 above, what will be the effect of the magnitude of the permissible emission on the actual emission? Trend analysis can be used to test the effects of one or more components of the trend.

Each of the components is tested using specific comparisons. The linear component of the trend test whether there is an overall increase (or decrease) in the dependent variable, as the independent variable increases or decreases.

In this case, the independent variable, which is the permissible emission, is fixated and it is expressed as the horizontal line (red line). Figure 1 shows that the actual emissions are consistent expect for the significant rise around February of 1999. The effect is more pronounced in the rise of EU1 (blue line) compare to EU2, (green line) with each surpassing the permissible emission respectively.

The test of the linear component of trend is to test the extent of the sharp increase and whether these respective sharp increases are significantly related. Looking at Figure 2, the significant is more distinct with the histogram. Figure 2 expresses the relationship of the AOR data of mercury emissions from EU1, EU2 and combined EUs, as well as the total permissible emission at the facility. The amount permissible per EU is half of the facility's total because there are only two emission units.

Figure 2 confirms the reflections express by Figure 1. Superimposing Figure 1 over Figure 2, there is a sharp increase in February 1994 but not enough to exceed the allowable emission per source unit, and this is more evident in EU1. In about February 1998 through 1999, there was a significant increase, which is also more evident in EU1 and enough to exceed the permissible emission per source unit.

Now back to Figure 1, a test of the linear component of the trend is to show the significant effect of the increase in actual emissions. If these were linear relationships between the permissible emission and the actual emission, then no components of trend other than the linear will be present. This is not the case. The slopes of the functions remain roughly steady for a period with significant spikes as the magnitude of the actual emissions on occasions, as the permissible emission remains constant.

The quadratic component of trend is used to test if the slope increases or decreases as permissible emission increases or decreases, which is not the case because the permissible emission remains constant. The cubic components test whether the slope changes twice (decreasing and then increasing or increasing and then decreasing) as the permissible emission increases or decreases. Again, just as with the quadratic component, the permissible emission is unchanging.

Trend analysis is computed as a set of orthogonal comparison using a particular set of coefficients¹. Coefficients for the linear, quadratic, and cubic components are given below:

2 means	3 means	· 4 means
Lin -1 1	Lin -1 0 1	Lin -3 -1 1 3
	Quad 1 -2 1	Quad 1 -1 -1 1
		Cubic -1 3 -3 1

5 means	<u>6 means</u>	<u>7 means</u>
Lin -2 -1 0 1 2	Lin -5 -3 -1 1 3 5	Lin -3 -2 -1 0 1 2 3
Quad 2 -1 -1 -1 2	Quad 5 -1 -4 -4 -1 5	Quad 5 0 -3 -4 -3 0 5
Cubic -1 2 0 -2 1	Cubic -5 7 4 -4 -7 5	Cubic -1 1 1 0 -1 -1 1

8 means	9 means	10 means∑
Lin -7 -5 -3 -1 1 3 5 7	Lin -4 3 -2 -1 0 1 2 3 4	Lin -9 -7 -5 -3 -1 1 3 5 7 9
Quad 7 1 -3 -5 -5 -3 1 7	Quad 28 7 -8 -17 -20 -17 -8 7 28	Quad 6 2 -1 -3 -4 -4 -3 -1 2 6
Cubic -7 5 7 3 -3 -7 -5 7	Cubic -14 7 13 9 0 -9 -13 -7 14	Cubic -42 14 35 31 12 -12 -31 -35 -14 42

From the ANOVA of the EU1, EU2 together and allowable emission per unit, the MSE (mean square error) is found to be 2.145 for the magnitude of the actual emissions against permissible emission. There are five groups with varying counts as shown in Table 1 below. For five groups from the illustrations above, the coefficient for the linear components will be:

2, -1, 0, 1, 2

Applying the comparison formula²,

TABLE 1: EMISSION TESTS DATA

t = L/S_L Where L =
$$\sum M_i a_i$$
 and S_L = $\sqrt{\sum \{(a_i^2/n_i) \text{ MSE}\}}$

 a_i is the coefficient applied to the i^{th} mean; n_i is the sample size of the i^{th} group; M_i is the i^{th} mean MSE = SSE/dfe and is obtained from ANOVA; The "t" is based on dfe = N - a degrees of freedom where "N" is the total number of subjects and "a" is the number of groups

¹ A coefficient is a constant used to multiply another value. For instance, in the linear transformation of Y = 3X + 7, the coefficient "3" is multiplied by the variable X. In the Linear combination of means, it follows that: $L = (2)M_1 + (-1)M_2 + (-1)M_3$ the three numbers in parentheses are coefficients.

² The method of computing planned comparisons among means is generalized as:

_	U1 CTUAL	TEST ALLOWED PER EU	EU2 ACTUAL	TEST ALLOWED PER EU		TOTAL ALLOWED
	0.1297	0.:	2	0.2	0.1297	0.4
	0.0559	0.:	2	0.2	0.0559	0.4
	0.00998	0.:	2 0.0027	7 0.2	0.01268	0.4
	0.00103	0.:	2	0.2	0.00103	0.4
	0.0241	O.:	2 0.024	1 0.2	0.0482	0.4
	0.0024	J 0.:	2 0.0024	0.2	0.0048	0.4
	0.0043	0. :	2 0.0478	0.2	0.0521	0.4
	0.06146	o.:	2 0.0027	7 0.2	0.06416	0.4
	0.0031	O.:	2 0.0028	3 0.2	0.0059	0.4
	0.0608	0.:	2 0.024	0.2	0.0848	0.4
	0.5771	l 0.:	2	0.2	0.5771	0.4
		0.:	2 0.010	0.2	0.016	0.4
	8.5538	0.:	2 0.737°	1 0.2	9.2909	0.4
MEAN	0.729513	3 0.	2 0.06612	3 0. 2	0.795636154	0.4

COUNT 12 13 9 13 13

 $L = M_i a_i$ therefore:

$$= (0.2)(-2) + (0.729513)(-1) + (0.066123)(0) + (0.4)(1) + (0.795636154)(2)$$
$$= (-0.4) + (-0.729513) + (0) + (0.4) + (1.5912722) = 0.8617592$$

 $S_L = \sqrt{\sum \{(a_i^2/n_i) \text{ MSE}\}}$ therefore:

$$= \sqrt{\sum ((-2)^2/12) + ((-1)^2/9) + ((0)^2/9) + ((1)^2/13) + ((2)^2/13)}$$
$$= \sqrt{0.829} = 0.9105$$

$$t = 0.8617592 / 0.9105 = 0.9465$$

The degrees of freedom for the t is equal to the degrees of freedom error in ANOVA which is:

$$N - a = 60 - 5 = 55$$

From the t-test for paired two samples for means (EU1 + EU2) the estimated probability value for two tails is 0.5872.

Therefore, the increase in the actual emissions is significant when compared to permissible emission even though the linear transformation was not observed..

Similarly, using the quadratic component of the trend, the increasing slope can be tested to see if the flattening out is significant. However, in this case, it is not possible to use the available data up to 1999, otherwise the quadratic components can be applied. The same condition applies to the cubic components.

	TOTAL					
TEST	ALLOWED					
ALLOW PER	AT					TOTAL
EU	FACILITY	YEAR	EU1	EU2	EU1 + EU2	ALLOWED*
0.2	0.4	1994	0.183	0.075	0.258	2
0.2	0.4	1995	0.0057	0.0071	0.0128	2
0.2	0.4	1996	0.006	0.0074	0.0134	2
. 0.2	0.4	1997	0.0172	0.0112	0.0284	2
0.2	0.4	1998	0.0342	0.0702	0.1044	2
0.2	0.4	1999	1.09	0.0069	1.0969	2

	TEST	_	TEST		TOTAL		
	ALLOW		ALLOW		ALLOWED		TOTAL
YEAR	PER EU	EU1	PER EU	EU2	AT FACILITY	EU1 + EU2	ALLOWED*
1994	0.2	0.183	0.2	0.075	0.4	0.258	2
1995	0.2	0.0057	0.2	0.0071	0.4	0.0128	2
1996	0.2	0.006	0.2	0.0074	0.4	0.0134	2
1997	0.2	0.0172	0.2	0.0112	0.4	0.0284	2
1998	0.2	0.0342	0.2	0.0702	0.4	0.1044	2
1999	0.2	1.09	0.2	0.0069	0.4	1.0969	2

	TEST			TOTAL	
	ALLOW			ALLOWED	
YEAR	PER EU	EU1	EU2	AT FACILITY	EU1 + EU2
1994	0.2	0.183	0.075	0.4	0.258
1995	0.2	0.0057	0.0071	0.4	0.0128
1996	0.2	0.006	0.0074	0.4	0.0134
1997	0.2	0.0172	0.0112	0.4	0.0284
1998	0.2	0.0342	0.0702	0.4	0.1044
1999	0.2	1.09	0.0069	0.4	1.0969

	1994-1999	AOR DATA	1
YEAR	EU1	EU2	EU1 + EU2
1994	0.183	0.075	0.258
1995	0.0057	0.0071	0.0128
1996	0.006	0.0074	0.0134
1997	0.0172	0.0112	0.0284
1998	0.0342	0.0702	0.1044
1999	1.09	0.0069	1.0969

TE				_		TOT ALLOWED A	T FACILITY	=111	=
ALLOW	PER EU	EU	1	E	U2	PER TEST		EU1 + EU2	
Mean Standard	0.2	Mean Standard	0.2226833	Mean Standard	0.029633333	Mean Standard	0.4	Mean Standard	0.252317
Error	1.3603E-09	Error	0.1756508		0.013616722	Error	2.7206E-09	Error	0.173188
Median	0.2	Median	0.0257	Median	0.0093	Median	0.4	Median	0.0664
Mode	0.2	Mode	#N/A	Mode	#N/A	Mode	0.4	Mode	#N/A
Standard		Standard		Standard		Standard		Standard	
Deviation	3.332E-09	Deviation	0.4302548	Deviation	0.03335402	Deviation	6.664E-09	Deviation	0.424222
Sample		Sample		Sample		Sample		Sample	
Variance	1.1102E-17	Variance	0.1851192	Variance	0.001112491	Variance	4.4409E-17	Variance	0.179964
Kurtosis	-3.3333333	Kurtosis	5.4611796	Kurtosis	-1.82207092	Kurtosis	-3.3333333	Kurtosis	4.953789
Skewness	1.36930639	Skewness	2.3224738	Skewness	0.967650442	Skewness	1.36930639	Skewness	2.202973
Range	0	Range	1.0843	Range	0.0681	Range	0	Range	1.0841
Minimum	0.2	Minimum	0.0057	Minimum	0.0069	Minimum	0.4	Minimum	0.0128
Maximum	0.2	Maximum	1.09	Maximum	0.075	Maximum	0.4	Maximum	1.0969
Sum	1.2	Sum	1.3361	Sum	0.1778	Sum	2.4	Sum	1.5139
Count	6	Count	6	Count	6	Count	6	Count	6
Largest(1)	0.2	Largest(1)	1.09	Largest(1)	0.075	Largest(1)	0.4	Largest(1)	1.0969
Smallest(1) Confidence	1	Smallest(1) Confidence	0.0057	Smallest(1) Confidence		Smallest(1) Confidence	1	Smallest(1) Confidence	0.0128
Level(95.0%)	3.4967E-09	Level(95.0%)	0.4515239	Level(95.0%)	0.03500284	Level(95.0%)	6.9934E-09	Level(95.0%)	0.445193

Assumed that e	ach of the popula	tions is normall	y distributed v	vith the same var	rlance (s2).	
Groups	Count	Sum	Average	Variance	•	
TEST			-		•	
ALLOW PER						
EU	6	1.2	0.2	1.11022E-17		
EU1	6	1.3361	0.2226833	0.185119154		
ANOVA						
Source of						
Variation	SS	df	MS	F	P-value	F crit
Between						
Groups	1 1					
(MSB)	0.0015436	1	0.0015436	0.016676835	0.899808271	4.964590516
Within]]					
Groups						•
(MSE)	0.92559577	10	0.0925596			
Total (TSS)	0.92713937	11				

Analysis of Variance is used to test hypothesis about differences between 2 or more means. When there are more than two means, it is possible to compare each mean with each other mean using t-tests. However, conducting multiple t-tests can lead to severe inflation. Analysis of variance can be used to test differences among several means for significance without increasing the Type I error rate.

LEGEND:

SS: Sum of squares;

df: Degree of Freedom = g-1 where G is the # of groups.

MS: Mean Square

F: Ratio of MSB to MSE i.e F=MSB/MSE & for null hypothesis to be true F will be ≤ 1

P-Value: Probability of a larger F

F Crit: From Statistical Tables used to test

ANOVA of different groups

Anova: Single F	actor - Emissic	on Unit 2 (I	EU2)
-----------------	-----------------	--------------	------

SUMMARY

Groups	Count	Sum	Average	Variance
TEST				
ALLOW PER				
EU	6	1.2	0.2	1.11022E-17
EU2	6	0.1778	0.0296333	0.001112491

ANOVA

Source of Variation	ss	df	MS	F	P-value	F crit
Between						
Groups Within	0.0870744	1	0.0870744	156.5395665	1.97094E-07	4.964590516
Groups	0.00556245	10	0.0005562			
Total	0.09263686	11				,

Fcrit of EU1 = Fcrit of EU1

However, it is shown from the results that ${\sf F}$ of EU1 is not equal to ${\sf F}$ of EU2.

Significant is the fact that while F of EU1 \leq 1; F of EU is considerably \geq 1

Anova: Single Factor for EU1, EU2 together

SUMMARY

Groups	Count	Sum	Average	Variance
TEST				
ALLOW PER				
EU	6	1.2	0.2	1.11022E-17
EU1	6			0.185119154
EU2	6	0.1778	0.0296333	0.001112491

ANOVA

Source of Variation	ss	df	MS	F	P-value	F crit
Between						
Groups						
(MSB)	0.13361527	2	0.0668076	1.076202234	0.365802673	3.682316674
Within						
Groups						
(MSE)	0.93115822	15	0.0620772			
· ·						
Total	1.0647735	17				

- 1. The results shows that both Fcrit of EU1 = Fcrit of EU2.
- 2. This confirms that though both units are within the same facility they are independent of each other and that any exceedances in emission, observes in one cannot be related to the other.
- 3. As a further proof that this is the case, F of both shows a significance and though the observed statistics is within the critical region, i.e. Fcrit for EU1 & EU2 together < Fcrit of EU1 and EU2 individually, the null hypothesis can be rejected because there is no indication of growth when dependent.
- 4. Accountability of any exceedances will have to be related to the specific unit.
- 5. F-test two-sample for variances also shows marked F critical of varying proportion. Here the independent variables are independently sampled with the allowance permitted for each unit or when both are consdered together.
- 6. The test shows that when an independent variable such as EU1 or EU2 appears to have an effect, it is very important to be able to state with confidence that the effect was really due to the variable and not just due to chance as in this case.

F-Test Two-Sample for Variances: EU1					
	TEST ALLOW PER EU	EU1			
Mean	0.2	0.222683333			
Variance	1.1102E-17	0.185119154			
Observations	6	6			
df	5	5			
F	5.9973E-17				
P(F<=f)					
one-tail	0				
F Critical					
one-tail	0.19800694				

F-Test Two-Sample for Variances: EU2					
	TEST ALLOW PER EU	EU2			
Mean	0.2	0.029633333			
Variance	1.11022E-17	0.001112491			
Observations	6	6			
df	5	5			
F	9.97962E-15				
P(F<=f)					
one-tail	0				
F Critical					
one-tail	0.198006944				

Page 6

F-Test Two-Sample for Variances EU1 + EU2					
	TOTAL				
	ALLOWED				
	AT FACILITY	EU1 + EU2			
Mann		0.252317			
Mean					
Variance	4.44E-17	0.179964			
Observations	6	6			
df	5	5			
F	2.47E-16				
P(F<=f) one-tai	0				
F Critical one-	0.198007				

t-Test: Paired Two Sample for Means (EU1)				
	TEST ALLOW PER EU	EU1		
Mean	0.2	0.222683333		
Variance	1.1102E-17	0.185119154		
Observations	6	6		
Pearson				
Correlation	, 0			
Hypothesized Mean				
Difference	0			
df	5			
t Stat P(T<=t)	-0.1291388			
one-tail t Critical	0.45114089			
one-tail P(T<=t)	2.01504918			
two-tail t Critical	0.90228179			
two-tail	2.57057764			

	t-Test: Paired Two Sample for Means (EU2)				
•	TEST ALLOW PER EU	EU2			
Mean	0.2	0.029633333			
Variance	1.11022E-17	0.001112491			
Observations	6	6			
Pearson					
Correlation	7.20952E-09				
Hypothesized Mean Difference	0				
df	5				
t Stat P(T<=t)	12.5115773				
one-tail t Critical	2.8936E-05				
one-tail P(T<=t)	2.015049176				
two-tail t Critical	5.7872E-05				
two-tail	2.570577635				

t-Test: Paired Two Sample						
for Mea	for Means (EU1 + EU2)					
	TOTAL					
	TOTAL ALLOWED					
	ALLOWED					
	FACILITY	EU1 + EU2				
Mean	0.4	0.252317				
	4.44E-17					
Variance		0.179964				
Observations	6	6				
Pearson						
Correlation	9.07E-09					
l						
Hypothesized						
Mean		· '				
Difference	0					
df	5					
t Stat	0.852735					
P(T<=t)						
one-tail	0.216373					
t Critical						
one-tail	2.015049					
P(T<=t)						
two-tail	0.432745	•				
t Critical						
two-tail	2.570578					

The t rather than the z (normal) distribution is used because the standard error has to be estimated from the data. If the hypothesis is true, t will have a t-distribution with 5 degrees of freedom. The t critical value for one-tail and two-tail are in the critical region when because of the values of P one-tail respectively. This provides a good statistical evidence to the hypothesis that the effects of the emission exceedance from any one unit can be better managed as an independent unit within the facility, and that it will not be a sound compliance practice to target the facility as a whole for non compliancee, rather it should be the independent unit responsible for the exceedance. The following tests can also be perform, and this will not change the statistical inference: t-test: 2 samples assuming equal variances, t-test: 2 samples for mean.

Groups	Count	Sum	Average	Variance		
TEST						
ALLOW PER						
EU	6	1.2	0.2	1.11022E-17		
EU1	6	1.3361	0.2226833	0.185119154		
EU2	6	0.1778	0.0296333	0.001112491		
TOTAL						
TOTAL ALLOWED						
ALLOWED AT FACILITY	6	2.4	0.4	4.44089E-17		
EU1 + EU2	6		0.4			
ANOVA						
ANOVA Source of						
	ss	df	MS	F	P-value	F crit
Source of	SS	df	MS	F	P-value	F crit
Source of Variation Between Groups		df	MS	F	P-value	F crit
Source of Variation Between Groups (MSB)	ss 0.42052045		<i>MS</i> 0.1051301	·	<i>P-value</i> 0.251591909	
Source of Variation Between Groups (MSB) Within				·		
Source of Variation Between Groups (MSB) Within Groups	0.42052045	4	0.1051301	·		
Source of Variation Between Groups (MSB) Within		4		·		
Source of Variation Between Groups (MSB) Within Groups	0.42052045	4	0.1051301	·		

EMISSION D	DATA (1994	- 199	9)				Conversi	on To PH
AIRS_ID	EU_NO		TEST_	TEST_ALLO	ACTUAL	UNIT	TEST_ALLOW	ACTUAL
0690046		2	30-Jun-95	0.0003	0.000004	02	0.1996	0.0027
0690046		2	01-Feb-96	70	15.4	20	0.2	0.0241
0690046		2	01-Feb-96	0.00034	0.000007	22	0.2	0.0024
0690046		2	10-Jan-97	70	32.2	28	0.2	0.0478
0690046		2	10-Jan-97	0.00034	0.000013	22	0.2	0.0027
0690046		2	10-Jan-97	999990	0.00279	PH	0.2	0.0028
0690046		2	10-Jan-97	999990	0.024	PH	0.2	0.024
0690046		2	21-Apr-98	0.0003	0.000007	02	0.2	0.016
0690046		2	29-Jan-99	70	258	28	0.2	0.7371
0690046		1	27-Jan-98	80	89	13	0.2	0.0024
0690046		1	23-Apr-98	80	89.6	13	0.2	0.0024
0690046		1	04-Feb-94	0.0003	0.000195	02	0.2	0.1297
0690046		1	26-Jul-94	0.0003	0.000084	02	0.2	0.0559
0690046		1	21-Jun-95	0.00034	0.000015	02	0.2	0.00998
0690046		1	30-Jan-96	0.00034	0.000005	22	0.2	0.00103
0690046		1	10-Jan-97	999990	0.0043	PH	0.2	0.0043
0690046		1	10-Jan-97	70	41.4	28	0.2	0.06146
0690046		1	10-Jan-97	0.00034	0.000018	22	0.2	0.0031
0690046		1	10-Jan-97	999990	0.0608	PH	0.2	0.0608
0690046		1	28-Jan-98	70	202		0.2	
0690046		1	29-Jan-99	70	2994		0.2	

	=		-		TOTAL
TEST DATE	TEST ALLOWEU1	ACTUAL	TEST ALLOWEU2	2 ACTUAL	ALLOWED
Feb-94	0.2	0.1297	0.2	0	0.4
Jul-94	0.2	0.0559	0.2	0	0.4
Jun-95	0.2	0.00998	0.2	0.0027	0.4
Jan-96	0.2	0.00103	0.2	0	0.4
Feb-96	0.2	0.0241	0.2	0.0241	0.4
Feb-96	0.2	0.0024	0.2	0.0024	0.4
Jan-97	0.2	0.0043	0.2	0.0478	0.4
Jan-97	0.2	0.06146	0.2	0.0027	0.4
Jan-97	0.2	0.0031	0.2	0.0028	0.4
Jan-97	0.2	0.0608	0.2	0.024	0.4
Jan-98	0.2	0.5771	0.2	0	0.4
Apr-98	0.2	0	0.2	0.016	0.4
Jan-99	0.2	8.5538	0.2	0.7371	0.4

	EMISSION TEST DATA						AOR DATA					
				TOTAL								TOTAL
TEST DATE	TEST ALLOWEU1	ACTUAL	EU2 ACTUAL	ALLOWED		EU TOTAL	YEAR	EU1		EU2	4.4	ALLOWED*
Feb-94	0.2	0.1297		(0.4	0.1297		1994	0.183		0.075	2
Jul-94	0.2	0.0559		(0.4	0.0559		1995	0.0057		0.0071	2
Jun-95	0.2	0.00998	0.0027	(0.4	0.01268		1996	0.006		0.0074	2
Jan-96	0.2	0.00103		(0.4	0.00103	1	1997	0.0172		0.0112	2
Feb-96	0.2	0.0241	0.0241	(0.4	0.0482		1998	0.0342		0.0702	2
Feb-96	0.2	0.0024	0.0024		0.4	0.0048		1999	1.09		0.0069	2
Jan-97	0.2	0.0043	0.0478	(0.4	0.0521						
Jan-97	0.2	0.06146	0.0027	(0.4	0.06416						
Jan-97	0.2	0.0031	0.0028	(0.4	0.0059						
Jan-97	0.2	0.0608	0.024	(0.4	0.0848						
Jan-98	0.2	0.5771		(0.4	0.5771						
Apr-98	0.2		0.016	(0.4	0.016						
Jan-99	0.2	8.5538	0.7371		0.4	9.2909						

	SUMMA	RY OF THE AN	IALYSIS AND	MERCURY EM		OR OGDEN MA	KTIN, LAKE (COUNTY	
EU1 A	CTUAL	EU2 AC	TUAL	TEST ALLO	W PER EU	EU TO	TAL	TOTAL AL	LOWED
Mean Standard	0.790305833	Mean Standard	0.095511111	Mean Standard	0.2	Mean	35274.3449	Mean	0.4
Error	0.707287153	Error	0.080358648	Error	1.19305E-09	Standard Error	143.9584618	Standard Error	2.3861E-09
Median	0.04	Median	0.016	Median	0.2	Median	35431.0031	Median	0.4
Mode Standard	#N/A	Mode Standard	0.0027	Mode Standard	0.2	Mode	#N/A	Mode	0.4
Deviation Sample	2.450114567	Deviation Sample	0.241075945	Deviation Sample	4.30159E-09	Standard Devia	519.0496156	Standard Devia	8.60319E-09
Variance	6.003061394	Variance	0.058117611	Variance	1.85037E-17	Sample Variand	269412.5035	Sample Variand	7.40149E-17
Kurtosis	11.86558233	Kurtosis	8.897270279	Kurtosis		Kurtosis	-0.30650034		#DIV/0!
Skewness	3.438317847	Skewness	2.977321476	Skewness	#DIV/0!	Skewness	-0.12499529	Skewness	#DIV/0!
Range	8.55277	Range	0.7347	Range		Range	1803.4241	Range	0
Minimum	0.00103	Minimum		Minimum		Minimum	34366.1297	Minimum	0.4
Maximum	8.5538	Maximum	0.7371	Maximum	0.2	Maximum	36169.5538	Maximum	0.4
Sum	9.48367	Sum	0.8596	Sum	2.6	Sum	458566.4837	Sum	5.2
Count	12	Count	9	Count	13	Count	13	Count	13
Largest(1)	8.5538	Largest(1)	0.7371	Largest(1)	0.2	Largest(1)	36169.5538	Largest(1)	0.4
Smallest (1)	0.00103	Smallest(1)		Smallest(1)		Smallest(1)	34366.1297	Smallest(1)	0.4
Confidence		Confidence		Confidence		` '		` '	
Levei(95.0%)	1.556729314	Level(95.0%)	0.185307495	Level(95.0%)	2.59943E-09	Confidence Lev	313.6585382	Confidence Lev	5.19886E-09

SUMMARY

Groups	Count	Sum	Average	Variance	
TEST					
ALLOW PER					
EU	13	2.6	0.2	1.85037E-17	
EU1	12	9.48367	0.790305833	6.003061394	

ANOVA

Source of	_					
Variation	SS	df	MS	F	P-value	F crit
Between						
Groups						
(MSB)	2.174396496	1	2.174396496	0.757357805	0.393149101	4.279343102
	,					
Within						
Groups (MSE)	66.03367533	23	2.871029362			
Total (TSS)	68.20807183	24				

Anova: Sin	gle Factor	- Emission	Unit 2	(EU2)

SUMMARY

Groups	Count	Sum	Average	Variance	
TEST					
ALLOW PER					
EU	13	2.6	0.2	1.85037E-17	
EU2	9	0.8596	0.095511111	0.058117611	

ANOVA

Source of Variation	SS	df	MS	F	P-value	F crit
Between						
Groups Within	0.058063526	1	0.058063526	2.497673448	0.129700792	4.351250027
Groups	0.464940889	20	0.023247044			
Total	0.523004415	21				

Anova: Sing	le Factor fo	r EU1, EU2 t	ogether			
SUMMARY						
Groups	Count	Sum	Average	Variance		
TEST						
ALLOW PER						
EU	· 13	2.6	0.2	1.85037E-17		
EU1	12	9.48367	0.790305833	6.003061394		
EU2	9	0.8596	0.095511111	0.058117611		
ANOVA Source of						·
Variation	SS	df	MS	F	P-value	F crit
Between				_		
Groups (MSB)	3.169799947	2	1.584899974	0.738840926	0.485890485	3.304819529
Within Groups (MSE)	66.49861622	31	2.145116652			

69.66841617

Total

Anova: Single Factor - EU1, EU2, EU1 + EU2	

SUMMARY

Groups	Count	Sum	Average	Variance
TEST				
ALLOW PER				
EU	13	2.6	0.2	1.85037E-17
EU1	12	9.48367	0.790305833	6.003061394
EU2	9	0.8596	0.095511 1 11	0.058117611
TOTAL				
ALLOWED AT				
FACILITY	13	5.2	0.4	7.40149E-17
EU1 + EU2	13	10.34327	0.795636154	6.538429621

ANOVA

Source of						
Variation	SS	df	MS	F	P-value	F crit
Between						
Groups						
(MSB)	4.881864615	4	1.220466154	0.463063909	0.762500442	2.539685795
· · ·						
Within						
Groups (MSE)	144.9597717	55	2.635632212			
Total	149.8416363	59				

F-Test Two-Sample for Variances:					
EU1					
	TEST				
	ALLOW PER				
	EU	EU1			
Mean	0.2	0.790305833			
Variance	1.85037E-17	6.003061394			
Observations	13	12			
df	12	11			
F	3.08238E-18				
P(F<=f)					
one-tail	0				
F Critical					
one-tail	0.368007846				

-	TEST	
	ALLOW PER	
	EU	EU2
Mean	0.2	0.095511111
Variance	1.85037E-17	0.058117611
Observations	13	9
df	12	.8
F	3.18384E-16	
P(F<=f)		
one-tail	0	
F Critical		
one-tail	0.351054297	

F-Test Two-Sample for Variances					
EU1 + EU2					
	TOTAL ALLOWED AT				
	FACILITY	EU1 + EU2			
Mean	0.4	0.795636154			
Variance	7.40149E-17	6.538429621			
Observations	13	13			
df	12	12			
F	1.132E-17				
P(F<=f)					
one-tail F Critical	0				
one-tail	0.372212483				

t-Test: Paired Two Sample for Means (EU1)				
	TEST			
	ALLOW PER			
	EU	EU1		
Mean	0.2	0.729513077		
Variance	1.85037E-17	5.550851148		
Observations	13	13		
Pearson				
Correlation	4.67775E-09			
Hypothesized Mean				
Difference	0			
df	12			
t Stat	-0.810342423			
P(T<=t)				
one-tail	0.216760077			
t Critical				
one-tail	1.782286745			
P(T<=t)				
two-tail	0.433520154			
t Critical				
two-tail	2.178812792			

t-Test: Paired Two Sample for Means (EU2)				
	TEST			
	ALLOW PER			
	EU	EU2		
Mean	0.2	0.066123077		
Variance	1.85037E-17	0.040850237		
Observations	13	13		
Pearson				
Correlation	0			
Hypothesized Mean				
Difference	o			
df	12			
t Stat	2.388251772			
P(T<=t)				
one-tail	0.017122094			
t Critical				
one-tail	1.782286745			
P(T<=t) .				
two-tail	0.034244188			
t Critical				
two-tail	2.178812792			

t-Test: Paired Two Sample for Means (EU1 + EU2)						
	TOTAL ALLOWED AT FACILITY	EU1 + EU2				
Mean	0.4	0.795636154				
Variance	7.40149E-17	6.538429621				
Observations	13	13				
Pearson						
Correlation	0					
Hypothesized Mean Difference df t Stat P(T<=t) one-tail t Critical one-tail P(T<=t)	0 12 -0.55786732 0.293595944 1.782286745 0.587191888					
t Critical	0.00.101000					
two-tail	2.178812792					